ARMA models Part 2: Moving average models (MA)

Beáta Stehlíková

Time Series Analysis

Faculty of Mathematics, Physics and Informatics, Comenius University

ARMA models Part 2: Moving average models (MA) - p.1/20

V.

Moving average process of the first order - MA(1)

Data from previous lecture

Ben Vogelvang: **Econometrics. Theory and Applications with EViews.** Pearson Education Limited, 2005.

Chapter 14.7. - The Box-Jenkins Approach in Practice

- Monthy data, January 1960 September 2002
- $pcocoa_t$ cocoa prices, we take logarithms and because of stationarity we will work with differences



Time ARMA models Part 2: Moving average models (MA) - p.3/20

Data from previous lecture

• Estimated ACF:



Series: diff(logpcocoa)

• One significantly nonzero autocorrelation and the remaining ones are nearly zero

Example from the first lecture

• Let u_t be a white noise, we define

$$x_t = u_t + u_{t-1}$$

• We computed:

$$E[x_t] = 0, \quad Var[x_t] = 2\sigma^2$$
$$Cov[x_t, x_{t+\tau}] = \begin{cases} \sigma^2 & \text{pre } \tau = 1\\ 0 & \text{pre } \tau = 2, 3, \dots \end{cases}$$
$$Cor[x_t, x_{t+\tau}] = \begin{cases} 1/2 & \text{pre } \tau = 1\\ 0 & \text{pre } \tau = 2, 3, \dots \end{cases}$$

• ACF is zero for $\tau = 2, 3, \ldots$ - exactly the property which we need

• Let u_t be a white noise, then

$$x_t = \mu + u_t - \beta u_{t-1}$$

is called a moving average process of the first order - MA(1)

- Wold representation: $x_t = \mu + \sum_{j=0}^{\infty} \psi_j u_{t-j}$ MA(1) process: $\psi_0 = 1, \psi_1 = -\beta, \psi_j = 0$ for j = 2, 3, ...
- Moments and ACF:

$$\begin{split} E[x_t] &= \mu, \quad Var[x_t] = (1+\beta^2)\sigma^2 \\ Cov[x_t, x_{t+\tau}] &= \begin{cases} -\beta\sigma^2 & \text{for } \tau = 1 \\ 0 & \text{for } \tau = 2, 3, \dots \end{cases} \\ Cor[x_t, x_{t+\tau}] &= \begin{cases} -\frac{\beta}{1+\beta^2} & \text{for } \tau = 1 \\ -\frac{\beta}{1+\beta^2} & \text{for } \tau = 1 \\ RMA \mod els & \text{Part } 2: \operatorname{Moving average models (MA)} - p.6/20 \end{cases} \end{split}$$

1. Let u_t be a white noise with distribution N(0, 4), we define

$$x_t = u_t + \frac{1}{2}u_{t-1}$$

Then:
$$E[x_t] = 0$$
, $Var[x_t] = (1 + (1/2)^2) \times 4 = 5$
 $Cor[x_t, x_{t+\tau}] = \begin{cases} \frac{1/2}{1+1/4} = 2/5 & \text{for } \tau = 1\\ 0 & \text{for } \tau = 2, 3, ... \end{cases}$

2. Let u_t be a white noise with distribution N(0, 1), we define

$$y_t = u_t + 2u_{t-1}$$

Then:
$$E[y_t] = 0$$
, $Var[y_t] = (1+4) \times 1 = 5$
 $Cor[y_t, y_{t+\tau}] = \begin{cases} \frac{2}{1+4} = 2/5 & \text{for } \tau = 1 \\ 0 & \text{for } \tau = 2, 3, \\ 0 & \text{ARMA models Part 2: Moving average models (MA) - p.7/20} \end{cases}$

• Let us take an MA(1) process, i.e. ACF of the form

$$Cor[x_t, x_{t+\tau}] = \begin{cases} -\frac{\beta}{1+\beta^2} & \text{for } \tau = 1\\ 0 & \text{for } \tau = 2, 3, \dots \end{cases}$$

• Suppose now that we are given the value $\rho_1 = \rho(1)$ and we want to find the coefficient β , i.e.



ARMA models Part 2: Moving average models (MA) - p.8/20

• We have therefore the equation:

 $\rho_1 = -\frac{\beta}{1+\beta^2} \Rightarrow \beta^2 + \frac{1}{\rho_1}\beta + 1 = 0$



 \rightarrow two solutions β_1 , β_2 , they satisfy $\beta_1\beta_2 = 1$.

Processes

have

$$x_t = \mu + u_t - \beta u_{t-1}, \quad x_t = \mu + u_t - \frac{1}{\beta} u_{t-1}$$

the same ACF

If we want a unique parametrization, we need an additional condition.
 ARMA models Part 2: Moving average models (MA) - p.9/20

• We will try to write the process in $AR(\infty)$ form:

 $x_t = \hat{\mu} + u_t + \psi_1 x_{t-1} + \psi_2 x_{t-2} + \psi_3 x_{t-3} + \dots$

- if is it possible, the process is called invertible
- For MA(1) process:

$$x_t = \mu + (1 - \beta L)u_t$$

(1 - \beta L)^{-1}x_t = (1 - \beta L)^{-1}\mu + u_t

 $(1 - \beta L)^{-1}$ exists for $|\beta| < 1$, then

$$(1 + \beta L + \beta^2 L^2 + ...)x_t = \mu/(1 - \beta) + u_t$$

$$x_t + \beta x_{t-1} + \beta^2 x_{t-2} + \ldots = \mu/(1-\beta) + u_t$$

MA(1) - invertibility

- We obtained invertibility condition for MA(1) process: $|\beta| < 1$
- Another way how to express it:
 - \diamond we have a process $x_t = \mu + (1 \beta L)u_t$
 - ♦ root of the polynomial $1 \beta L$ is $1/\beta$
 - ♦ invertibility condition means that root of $1 - \beta L = 0$ has to be in absolute value greater than 1, i.e. outside of the unit circle

MA(1) - computation of PACF

• Recall the general formula:

(1)
$$\Phi_{kk} = \frac{\det \begin{pmatrix} 1 & \rho(1) & \dots & \rho(1) \\ \rho(1) & 1 & \dots & \rho(2) \\ & \dots & & & \\ \rho(k-1) & \rho(k-2) & \dots & \rho(k) \end{pmatrix}}{\det \begin{pmatrix} 1 & \rho(1) & \dots & \rho(k-1) \\ \rho(1) & 1 & \dots & \rho(k-2) \\ & \dots & & & \\ \rho(k-1) & \rho(k-2) & \dots & 1 \end{pmatrix}}$$

• For MA(1) we have $\rho(k) = 0$ for k = 2, 3, ...

MA(1) - computation of PACF

• PACF is not zero after some lags (as it holds for AR):

$$\begin{split} \Phi_{11} &= \rho(1) \\ \Phi_{22} &= \frac{\det \begin{pmatrix} 1 & \rho(1) \\ \rho(1) & \rho(2) \end{pmatrix}}{\det \begin{pmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{pmatrix}} = \frac{\det \begin{pmatrix} 1 & \rho(1) \\ \rho(1) & 0 \end{pmatrix}}{\det \begin{pmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{pmatrix}} = \frac{-\rho(1)^2}{1 - \rho(1)^2} \\ \Phi_{33} &= \frac{\det \begin{pmatrix} 1 & \rho(1) & \rho(1) \\ \rho(1) & 1 & \rho(2) \\ \rho(2) & \rho(1) & \rho(3) \end{pmatrix}}{\det \begin{pmatrix} 1 & \rho(1) & \rho(2) \\ \rho(1) & 1 & \rho(1) \\ \rho(2) & \rho(1) & 1 \end{pmatrix}} = \frac{\det \begin{pmatrix} 1 & \rho(1) & \rho(1) \\ \rho(1) & 1 & 0 \\ 0 & \rho(1) & 0 \end{pmatrix}}{\det \begin{pmatrix} 1 & \rho(1) & \rho(2) \\ \rho(1) & 1 & \rho(1) \\ 0 & \rho(1) & 1 \end{pmatrix}} = \frac{\rho(1)^3}{1 - 2\rho(1)^2} \\ \Phi_4 &= \frac{-\rho(1)^4}{(1 - \rho(1)^2)^2 - \rho(1)^2} \\ \cdots \end{split}$$

Real data - cocoa prices

- Data from the beginning of the lecture
- MA(1) model for differences of logarithms (variable *y* in the output from R):

```
> sarima(y,0,0,1,details=FALSE)
$fit
Call:
stats::arima(x = xdata, order = c(p, d, q), seasonal = list(order = c(P, D,
        Q), period = S), xreg = xmean, include.mean = FALSE, optim.control = lis$
        REPORT = 1, reltol = tol))
Coefficients:
            mal xmean
            0.3520 0.0024
s.e. 0.0402 0.0037
sigma^2 estimated as 0.003897: log likelihood = 693.61, aic = -1381.21
```

• ACF of residuals:

ACF of Residuals

• Ljung-Box statistics:



• Model is OK.

ARMA models Part 2: Moving average models (MA) - p.15/20

VI.

Moving average process of order q - MA(q)

• Let u_t be a white noise, then

$$x_{t} = \mu + u_{t} - \beta_{1} u_{t-1} - \beta_{2} u_{t-2} - \dots - \beta_{q} u_{t-q}$$

is called a moving average proces of q-th order - MA(q)

- Wold representation: $x_t = \mu + \sum_{j=0}^{\infty} \psi_j u_{t-j}$ MA(q) process: $\psi_0 = 1, \psi_1 = -\beta_1, \dots, \psi_q = -\beta_q, \psi_j = 0$ for $j > q \to MA(q)$ proces is always stationary
- Moments, ACF, PACF:

$$E[x_t] = \mu, \quad Var[x_t] = (1 + \beta_1^2 + \dots \beta_q^2)\sigma^2$$

 $Cov[x_t, x_{t+\tau}] = 0$ for $\tau = q + 1, q + 2, ...$

$$\Rightarrow Cor[x_t, x_{t+\tau}] = 0 \text{ for } \tau = q+1, q+2, \dots$$

. . .

• Computation of the first q autocorrelations (we can assume $\mu = 0$):

$$Cov[x_t, x_{t+\tau}] = E[(u_t - \beta_1 u_{t-1} - \dots - \beta_q u_{t-q}) \times (u_{t+\tau} - \beta_1 u_{t+\tau-1} - \dots - \beta_q u_{t+\tau-q})]$$

$$= E[u_t(u_{t+\tau} - \beta_1 u_{t+\tau-1} - \dots - \beta_q u_{t+\tau-q})] -\beta_1 E[u_{t-1}(u_{t+\tau} - \beta_1 u_{t+\tau-1} - \dots - \beta_q u_{t+\tau-q})]$$

$$-\beta_q E[u_{t-q}(u_{t+\tau} - \beta_1 u_{t+\tau-1} - \ldots - \beta_q u_{t+\tau-q})]$$

• Continued:

$$\tau = 1 \implies \gamma(1) = (-\beta_1 + \beta_1 \beta_2 + \ldots + \beta_{q-1} \beta_q) \sigma^2$$

$$\tau = 2 \implies \gamma(2) = (-\beta_2 + \beta_1 \beta_3 + \ldots + \beta_{q-2} \beta_q) \sigma^2$$

$$\cdots$$

$$\tau = q \implies \gamma(q) = (-\beta_q) \sigma^2$$

• **PACF** - substitution of ACF into (1)

• Invertibility:

$$x_t = \mu + u_t - \beta_1 u_{t-1} - \beta_2 u_{t-2} - \dots - \beta_q u_{t-q}$$

$$x_t = \mu + (1 - \beta_1 L - \dots - \beta_q L^q) u_t$$

• Existence of $(1 - \beta_1 L - \ldots - \beta_q L^q)^{-1}$ - roots of $1 - \beta_1 L - \ldots - \beta_q L^q = 0$ have to be outside of the unit circle